5th World Congress on Civil, Structural, and Environmental Engineering (CSEE'20) Lisbon, Portugal Virtual Conference – October, 2020 Paper No. ICEPTE 125 DOI: 10.11159/icepte20.125

Analysis of Global CO₂ Emissions: China's Impacts by Two Stage Least Squares and Threshold Model

Zhiguang Song¹

¹Sophia University 7-1 Kioi-cho, Chiyoda-ku, Tokyo, Japan, genv-co@sophia.ac.jp sosico.song@gmail.com

Abstract - As the "Going Out" strategy was rolled out after WTO accession, the vigorous booming of China's outward foreign direct investment (OFDI) and international trade have been reported to pose a significant global environmental issue. On that basis, the present study exploits the fixed effects-two stage least squares (FE-2SLS) to delve into China's impacts (e.g., its overseas investment and foreign trade) and economic and institutional distance with China on CO₂ emissions per capita of 146 countries & regions (2003 to 2016), which is achieved based on instrumental variables. Subsequently, the threshold regression model with fixed effects is adopted to capture how those China's overseas economic activities will impact global air pollution at various phases of economic or institutional distance by taking it as structural breaks or threshold regimes. Consequently, given sustainability and global emission reduction, first, assessed results of the FE-2SLS empirically support that China's overseas investment should proactively focus on high-value-added sectors, while the marginal amount in local polluting industries should be controlled. Moreover, China should continue to optimize the trade structure to make its imports more environmental-friendly. Second, as overall suggested from the tests of threshold effect and results of threshold regression with the geographic distribution of different threshold regimes, the Chinese government should enhance its government effectiveness and creditability of the legal system, as an attempt to enhance the political proximity with developed countries. Furthermore, Chinese multinationals are required to attempt to proactively develop more economic cooperation with developed markets in a high degree of administration and environmental standards; political claims in the mentioned process should be limited.

Keywords: CO2 emissions per capita; China; Two Stage Least Squares; Threshold Model; Political Proximity

1. Introduction

Since the implementation of "Going Out" strategy after WTO accession in the 2000s, China's outward foreign direct investment (OFDI) has skyrocketed, especially when China has ushered into a "New Normal" era, as accompanied with a series of global policies and economic strategies (e.g., the proposal of the One Belt & One Road initiative (BRI) and establishment of the Asian Infrastructure Investment Bank (AIIB) unveiled in the 2010s after the world economic and financial crisis). According to the Statistical Bulletin of China's Outward Foreign Direct Investment (2016), China's OFDI flows first exceeded the amount of inward foreign direct investment (IFDI) in 2015. Moreover, newly increased OFDI by Chinese domestic investors rapidly rose to \$1456.7 billion in 2015 from only \$28.5 billion in 2003, ranking second worldwide, making it a true investment country.

Moreover, as one of the critical sectors of economic growth, foreign trade in China has been overall undergoing steady growth, especially since its WTO accession in 2001. According to the China Statistical Yearbook (2016), China's total trade value monotonously increases until 2009 affected by the world economic crisis and weak global demand. Subsequently, as the world's largest trading nation since 2013 and largest bilateral trading partner for most countries, China's total trade value peaked at \$4301.53 billion in 2014, when China has ushered into the "New Normal" era with ongoing industrial upgrading.

2. Literature Reviews

Currently, existing literature on the relationship of investment facilitation & trade liberalization and environment suggests a potentially high level of interaction. To be specific, Grossman and Kruger (1995) first decomposed the impact mechanism of foreign direct investment (FDI) on the environment into three channels: scale effect, composition effect, and technique effect, which exert different or even reversed impacts on environment respectively [1]. Inspired by the mentioned

study, Feng (2008) determined the relationship between inflows of FDI and local air pollution in China, suggesting that FDI can facilitate a developing country's environment when the multinationals crowd out inefficient local firms, change the industrial composition, advance technology and boost productivity & energy efficiency [2]. Likewise, Eskeland and Harrison (2003) identified little support for the pollution haven hypothesis when the U.S. invests in four countries (Cote d'Ivoire, Mexico, Morocco, and Venezuela). Also, they reported that foreign plants are significantly more energy-efficient and use cleaner types of energy than domestic-owned plants [3]. At a larger scale, however, Grimes et al. (2003) analyzed 66 less developed countries, suggesting that foreign investment is more concentrated in those industries that require more energy and positively linked with local CO_2 emissions. Besides, transnational corporations may relocate highly polluting industries to delve into the relationships between FDI and CO_2 emissions in 112 countries at different income levels. It was reported that FDI inflows in middle-income countries will promote CO_2 emissions, whereas no such significant relationship was identified in high-income countries. In low-income countries, however, CO_2 emissions can even hinder the entry of FDI [5].

Given environmental pollution embodied in international trade, Shui and Harriss (2006) verified that 7%-14% of China's CO₂ emissions are caused by its exports to the United States [6]. Inversely, Wyckoff and Roop (1994) calculated the embodied carbon emissions in manufactured goods imported by the OECD countries, taking up nearly 13% of its overall emissions [7]. At a larger scale, Peters et al. (2008) exploited the input-output data of 2001 to delve into the embodied carbon emissions in international trade in 87 countries. It was reported that carbon emissions embodied in international trade took up 1/4 of the world's total emissions, with China's exports and imports carbon emissions respectively accounting for 24% and 7% of its domestic carbon emissions [8].

Moreover, concerning the role of institutional factors related to China in cross-border investment and trade. Kolstad and Wiig (2012) reported that institutional quality of host country exhibits a negative correlation with China's OFDI, which is attracted to large markets (OECD countries), and countries with a combination of considerable natural resources and poor institutions (non-OECD countries) [9]. Likewise, Cheung and Qian (2009) utilized China's OFDI data to Africa, indicating that corruption and legal defects in the region have positive impacts on China's OFDI [10]. Generally, Habib and Zubawicki (2002) proposed that global FDI is not likely to introduce under a severe level of corruption in the host country [11], while Aleksynska and Havrylchyk (2012) suggested that especially for the southern countries, the greater institutional distance will hinder the FDI inflow [12].

Thus, since the global climate change with increasing greenhouse gas (GHG) emissions is becoming a contentious issue worldwide, and when China's economy is developing in a broader and more advanced manner, with the emerging process of growing overseas investment and transitions in industrial and trade structure. China's OFDI and foreign trade will likely to pose global environmental issues other than economic benefits, as it is becoming the second-largest investor (inflows) and the largest industrializing & trading nation globally. Thus, in this emerging field rarely discussed by existing studies, the present study will first aim to demonstrate the link of impacts exerted by China (e.g., its overseas investment and foreign trade) and economic and institutional distances with China on CO₂ emissions per capita of 146 countries & regions (2003 to 2016), which is achieved based on the fixed effects-two stage least squares with instrumental variables controlling over endogeneity. Next, it is noteworthy that the threshold regression model with fixed effects is adopted to capture how those China's overseas economic activities will impact global air pollution, at various phases of economic or institutional distance by taking it as structural breaks or threshold regimes.

3. Threshold Regression Model

The threshold model was introduced by Tong (1978), Tong and Lim (1980), which are overall employed in discrete time series that exhibit piece-wise linearity [13]. Subsequently, the threshold models and other nonlinear time series models have led to the development and sometimes the re-discovery of numerous data analytic techniques [14]. Given the definition of the threshold model, coefficients in the model are inconstant and can be altered by the endogenously determined threshold variables. Furthermore, it allows to split the sample into different regimes, so it is not required to enter dummy variables based on unknown structural breaks [15].

A single threshold regression model with fixed effects is written as:

$$y_{it} = \mu_i + \theta' x'_{it} + \beta_1 x_{it} I(q_{it} \le \gamma) + \beta_2 x_{it} I(q_{it} \succ \gamma) + \varepsilon_{it}$$
(1)

In Eq. (1), i and t indicate area and time effect respectively, I (\cdot) represents an indicator function, it is equal to 1 (there exists a single-threshold effect) or 0 (there exists a linear effect). q_{it} denotes the threshold variable which refers to the economic or institutional distances in the present study. γ represents the threshold value that divides the equation into two regimes with different threshold coefficients of β_1 and β_2 . Besides, X_{it} denotes the core explanatory variable that respectively refers to China's overseas investment or imports respectively, and y_{it} represents the explained variable refers to CO_2 emissions per capita of 146 countries & regions. Lastly, X'_{it} in the present study is a set of control variables representing the rest of variables other than threshold variable that bring a linear relationship on explained variable with coefficient θ' , while μ_i represents the individual effect and ε_{it} indicates the disturbance term. Moreover, individual fixed effect (μ_i) will be eliminated via fixed effect transformation by deducting its own average value in the present study.

Commonly, the estimated threshold value $(\overline{\gamma})$ can be derived by minimizing the residual sum of squared errors $S_n(\gamma) = \mathcal{E}'_{it} \mathcal{E}_{it}$, and once the threshold value is obtained, the corresponding parameter estimates of threshold $(\overline{\beta})$ will be given by the ordinary least squares [16].

Here, when the test is performed on the significance of threshold effect at relevant confidence level, whether there exists a threshold effect is verified, which satisfies the purpose of verifying whether the coefficients are identical in the respective regime, i.e., the null hypothesis (H_0 : $\beta 1=\beta 2$) against the alternative hypothesis (H_1 : $\beta 1\neq\beta 2$). Accordingly, the null hypothesis (of linearity) can be rejected, a single threshold effect nonlinear regression will take place (with two threshold regimes). Besides, the F statistic under the null hypothesis is defined as:

$$F_1 = \frac{S_0 - S_1(\bar{\gamma})}{\bar{\sigma}^2} \tag{2}$$

Where the S₁ and S₀ respectively denote the residual sum of squared errors with and without threshold effect respectively, $\overline{\sigma^2} = \frac{1}{n} S_1(\overline{\gamma})$ indicates the variance of the residual. Given the null hypothesis of H_0 , the threshold value is not identified, and F₁ exhibits a nonstandard asymptotic distribution. Accordingly, Hansen (1999) developed a repetitive bootstrap approach on the critical values of the F statistic, as an attempt to make the statistic asymptotically distributed and verify the significance of the threshold effect by determining the asymptotic p-values [17]. In the present study, it will run 1000 bootstraps by Stata 15.

4. Data and Variables

Given the availability and representativeness of data, the variables applied in the present study packing in a panel set of 146 countries & regions (2003 to 2016) are listed in Table 1. Here, "CO₂ Emissions per capita" (ton/per, CO2P) of 146 countries & regions are represented as explained variable. Besides, "China's Outward Foreign Direct Investment Stock" (10 Thousand USD, OFDI) and "China's Import Value" (10 Thousand USD, IV) are set as core explanatory variables respectively. "Gap of Industrial Output" (10 Thousand USD, IG) and "Gap of Worldwide Governance Indicators" (WGIB) in absolute value are treated as threshold variable respectively to verify the relevant threshold effect in threshold model, and "GDP (Nominal) Per capita" (10 Thousand USD/Per, GDP) of 146 Countries & Regions is regarded as the control variable. As a supplement, "Ratio of OFDI to GDP" (%, OG), "Ratio of Import Value to GDP" (%, ITG) and "GDP Per capita" (10 Thousand USD/Per, GDPC) of China are adopted as instrumental variables conducted in fixed effects-two stage least squares. In addition, the present study will transform the variables into logarithms to reduce the fluctuation of the model without changing the assessment trend.

Variables	Abbreviation	Unit	Description		
Explained Variable	CO2P	Ton/per	CO ₂ Emissions Per capita		
Explanatory Variables	OFDI	10 Thousand USD	China's Outward Foreign Direct Investment Stock		
	IV	10 Thousand USD	China's Import Value		
	IG	10 Thousand USD	Gap of Industrial Output (Absolute Value)		
	WGIB	- Gap of Worldwide Governance Indicators (Abso			
	GDP	10 Thousand USD/per	GDP Per capita of 146 Countries & Regions		
Instrumental Variables	OG	%	Ratio of OFDI to GDP of China		
	ITG	%	Ratio of Import Value to GDP of China		
	GDPC	10 Thousand USD/per	GDP Per capita of China		

Table 1: Variable Definitions.

Source: Our World in Data; Statistical Bulletin of China's Outward Foreign Direct Investment; UN Comtrade; World Bank, and China Statistical Yearbook in 2003-2016.

5. Regression Analysis and Estimates

As a reference, the present study will simulate the linear regression before the threshold regression model, by applying the two stage least squares that considers the endogeneity problem, which arises when explanatory variables and the error term are correlated, or omitted variable is confounding both independent and dependent variables [18]. Accordingly, fixed effects-two stage least squares (FE-2SLS) will be utilized in the present study by employing exogenous instrumental variables to perform the endogenous bias correction, based on various of corresponding tests and results in Table 2.

Table 2: Instrumental Variables Estimation of the Multiple Regression Model.

	OLS	FE	2SLS	FE-2SLS			
lnOFDI	-0.066***	0.008**	0.379*	0.113***			
lnIV	0.071***	0.017***	-0.479**	-0.208***			
lnIG	-0.222***	-0.057***	-0.827**	-0.094*			
lnWGIB	-0.157***	0.002	-0.212***	0.031*			
lnGDP	0.927***	0.238***	1.367***	0.394***			
Intercept	-5.84e-07	-1.41e-07	-4.06e-06	-8.46e-07			
Hausman Test for FE vs RE	505.810***						
Hausman Test for Endogeneity		13.370**					
Durbin-Wu-Hausman Test			36.843***				
Underidentification Test							
Anderson canon. corr. LM statistic			6.486	18.832			
(without robust option)			(0.039)	(0.000)			
Kleibergen-Paap rk LM statistic			7.132	17.230			
(with robust option)			(0.028)	(0.000)			
Overidentification Test							
Sargan statistic			0.457	0.007			
(without robust option)			(0.499)	(0.935)			
Hansen J statistic			0.402	0.005			
(with robust option)			(0.526)	(0.943)			
R ²	0.777	0.178					
Observations	2044	2044	2044	2044			

Note: 1) ***, **, * denote significance at 1%, 5%, and 10% levels respectively, and P-value also in parentheses. 2) OLS: Ordinary Least Squares, FE: Fixed Effects, RE: Random Effects, 2SLS: Two Stage Least-Squares, FE-2SLS: Fixed Effects-Two Stage Least-Squares.

To be specific, when OG (Ratio of OFDI to GDP), ITG (Ratio of Import Value to GDP) and GDPC (GDP Per capita) of China are taken as instrumental variables for OFDI and IV considering endogenous explanatory variables, the null hypothesis of Hausman test for endogeneity (H_0 : all explanatory variables are exogenous) and Durbin-Wu-Hausman (DWH) test (H_0 : endogenous regressors are in fact exogenous) are both rejected in a statistically significant manner, demonstrating that OFDI and IV are indeed endogenous explanatory variables in the present study. Subsequently, the implementation of the instrumental variables will be verified by the under-identification test on the correlation between instrumental variables and endogenous explanatory variable (H_0 : instrumental variables are not correlated with endogenous explanatory variable) and by the over-identification test on the exogeneity of instrumental variables (H_0 : instrumental variables are not correlated with endogenous explanatory variable) with disturbance term), which are supported to be valid.

Accordingly, how the assessed results of the FE-2SLS controlling endogenous bias differ from the results of the FE in Table 2 suggest that the OFDI distinctly exhibits a higher positive coefficient (0.113) significantly. It is therefore revealed that instrumental variables of the OG, ITG and GDPC, representing China's development of the overseas economic cooperation, tend to positively impact the global level of CO_2 emission. In contrast, IV in FE-2SLS turns into opposite and negative coefficient (-0.208) significantly compared with the result in FE, which can be interpreted as a CO_2 mitigation effect concealed in China's global imports when China's economy is transforming to a more open and mature market.

On that basis, the threshold regression model with fixed effects will be simulated to capture how China's OFDI and imports will impact air pollution of 146 countries & regions, under every phase of economic or institutional distance treated as threshold regimes. Simultaneously, given the lagged effects of overseas investment and international trade, along with the correction for the endogenous problem on some level, assessment of the threshold regression model in the present study will apply one-year lagged OFDI and IV as core explanatory variables.

First, the threshold effect test will be performed in Table 3, separated in the relationship of L.OFDI (One year lag OFDI) and CO2P, as well as that of L.IV (One-year lag IV) and CO2P. The first column lists the threshold variables holding a threshold effect that exhibits statistical significance at the 5% and 1% level out of the explanatory variables in Table 1. Thus, the third column respectively lists the threshold values dividing the model into two or three regimes; the last two columns represent F statistic with its probability value (p-value) and 95% confidence interval. On that basis, the test results reveal that the threshold variable of WGIB exhibits a single threshold effect of 1.179 at a significant level of 1% in the relationship of L.OFDI and CO2P; thus, the sample is split into two regimes. Likewise, GDP holds a double threshold effect with a value of 0.065 and 0.851 (10 Thousand USD/Per) successively, cutting the sample into three regimes. Moreover, in the relationship between L.IV and CO2P, only WGIB exhibits a significantly single threshold effect.

Explanatory Variable: L.OFDI*							
Threshold Variable	Threshold Effect	Threshold Value	F statistic	95% Confidence Interval			
WGIB	single threshold	1.179	210.62***	[1.159,1.194]			
GDP	single threshold	0.065	248.77***	[0.064,0.066]			
	double threshold	0.851	142.51***	[0.820,0.860]			
Explanatory Variable: L.IV*							
Threshold Variable	Threshold Effect	Threshold Value	F statistic	95% Confidence Interval			
WGIB	single threshold	1.107	84.99***	[1.094,1.136]			

Table 3: Threshold Effect Test and Threshold Estimates.

Note: 1) ***, **, * denote significance at 1%, 5%, and 10% levels respectively. 2) L.OFDI: lagged OFDI, L.IV: lagged IV.

Furthermore, assessed results of the threshold regression model with fixed effects are listed in Table 4, suggesting that: 1) At higher regime of the threshold regime 1, 2 and 3 (low, mid and high, respectively), marginal growth in L.OFDI or L.IV will cut the CO2P in a statistically significant manner. For instance, at the high regime of WGIB (>1.179) in the relationship of L.OFDI and CO2P, 1% increase in one-year lagged OFDI of China is demonstrated to lower 0.043% CO₂ emissions per capita of local country & region, of which the emissions level is noticeably less than that in the low regime. 2) When adding the cross-term combination of year dummy and OFDI or IV (omitted in the table) that is adopted to capture the time fixed effects (e.g., impact of policies or shocks that is restricted to a set period). In the relationship of L.OFDI and CO2P, OFDI in 2014-2016 exerts an exacerbated effect on CO2P, while IV in 2007-2008 exhibits the identical effect in the relationship of L.IV and CO2P.

Explanatory	Threshold Variable		Explanatory	Threshold Variable
Variable: L.OFDI	lnWGIB	lnGDP	Variable: L.IV	lnWGIB
lnIV	0.024***	0.015***	lnOFDI	0.007*
lnIG	-0.039***	-0.029**	lnIG	-0.033**
lnWGIB	0.002	0.001	lnWGIB	-0.003
lnGDP	0.150***	0.129***	lnGDP	0.201***
Threshold Regime 1	0.021***	0.093***	Threshold Regime 1	0.027***
Threshold Regime 2	-0.043***	0.019***	Threshold Regime 2	-0.038***
Threshold Regime 3		-0.033***		
Intercept	-0.005	0.003	Intercept	0.031***
\mathbb{R}^2	0.258	0.322	\mathbb{R}^2	0.223
Observations	1898	1898	Observations	1898

Table 4: Threshold Regression Model Estimates.

Note: 1) ***, **, * denote significance at 1%, 5%, and 10% levels respectively. 2) Threshold Regime 1, 2 and 3 indicate low, mid, and high regimes respectively.

6. Verification of Threshold Variable

Last, to capture the spatial variability among 146 countries & regions (2003 to 2016), threshold variable of WGIB is added into the ArcGis 10 to plot a distribution diagram in 2016, which is statistically significant in the test of threshold effect under both relationships between L.OFDI (L.IV) and CO2P. According to Fig. 1 combined with the assessed results of Table 4, at the low regime of WGIB (\leq =1.107 or \leq =1.179), i.e., when the institutional distance between China and the specific country & region is less than 1.107 or 1.179, the marginal increase in China's investment or imports is demonstrated to produce 0.021% or 0.027% CO₂ emissions per capita of local areas in Asia, Africa, East Europe and South America on the whole. In contrast, at the high regime of WGIB (>1.107 or >1.179) that connects a greater institutional distance, it exerts a mitigation effect and leads to the decrease in the emission levels by 0.043% or 0.038%, when China increases its overseas investment or imports, which are intensively distributed in West & North Europe, North America, Oceania, and Northeast Asia.



Fig. 1: Spatial Distributions of Threshold Variable (WGIB) at 2016.

7. Conclusions & Suggestions

Thus far, the present study first aims to demonstrate the link between China's rising impacts (e.g., its overseas investment and foreign trade) and economic and institutional distance with China, and CO_2 emissions per capita of 146 countries & regions (2003 to 2016). On that basis, this study aims to analyze how China's global economic activities will impact global CO_2 emissions by exploiting the fixed effects-two stage least squares (FE-2SLS) with instrumental variables regulating endogeneity. As revealed from the assessed results, increasing overseas investment from China over the years is demonstrated to exacerbate climate impacts, while China's imports is evidently demonstrated to reduce global CO_2 emissions in the present study.

Thus, given sustainability and global emissions reduction, when China's economic development is growing openly and advanced, China's capital investment in the future should proactively focus on high-value-added sectors (e.g., information technology, financial and service industries) in developed and mature markets, while the marginal amount in polluted industries (e.g., manufacturing, construction and primary energy-related industries) should be controlled. Moreover, China should place the optimization and promotion in trade structure as one of the top policy priorities, progressively eliminate its low value-added & high pollution sectors, and develop Hi-tech industries to make its imports more environmental-friendly.

Next, the threshold regression model with fixed effects is utilized to capture how overseas investment and imports of China will impact global CO_2 emissions, at various phases of economic or institutional distance by taking it as structural breaks or threshold regimes. As a result, the marginal increase in one-year lagged OFDI or IV of China will reduce the level of global CO_2 emissions significantly at a higher threshold regime. Besides, cross-terms combined of year dummy and OFDI display significant positive correlations with CO2P in 2014-2016; cross-terms of year dummy and IV have intensified effect on CO2P in 2007-2008. Lastly, countries & regions in the low regime of WGIB (<=1.107 or <=1.179) that connect with higher CO_2 emissions are primarily located in Asia, Africa, East Europe and South America. However, the mentioned areas in the high regime of WGIB (>1.107 or >1.179) that exert a mitigation effect on emission levels are intensively distributed in West & North Europe, North America, Oceania and Northeast Asia.

In brief, the Worldwide Governance Indicators (WGI) is determined to express the degree of the political, administrative and legal system from the World Bank. To be specific, China is empirically likely to cut environmental effect on relatively developed markets with lower political proximity; besides, it can exert an aggravated effect on developing countries that exhibit relatively higher political proximity on the whole in overseas investment and foreign trade. Thus, given sustainability and global emission reduction, the Chinese government should enhance its government effectiveness and creditability of the legal system, expedite market reform and boost greater openness & transparency, as an attempt to facilitate the political proximity with developed countries. Furthermore, compared with the period of rapid economic growth in the 2000s primarily dependent of the growth of foreign trade, Chinese multinationals should press ahead of the "Going Out" and attempt to proactively develop more economic cooperation with developed markets complying with a high degree of administration and environmental standards; moreover, political claims in the mentioned process should be limited.

References

- [1] G. Grossman and A. Krueger, "Economic growth and the environment," *The Quarterly Journal of Economics*, vol. 110, issue 2, 1995, pp. 353-377.
- [2] L. Feng, "Does foreign direct investment harm the host country's environment? Evidence from China," *SSRN Electronic Journal*, 2008.
- [3] G. Eskeland and A. Harrison, "Moving to greener pastures? Multinationals and the pollution haven hypothesis," *Journal of Development Economics*, vol. 70, issue 1, 2003, pp. 1-23.
- [4] P. Grimes, J. Rey, K. Grimes, and J. Kentor, "Exporting the greenhouse: Foreign capital penetration and CO₂ emissions 1980–1996," *Journal of World-Systems Research*, vol. 9, issue 2, 2003, pp. 261.
- [5] R. Hoffmann, L. C. Ging, B. Ramasamy and M. Yeung, "FDI and pollution: A Granger causality test using panel data," *Journal of International Development*, 2005.
- [6] B. Shui and R. Harriss, "The role of CO₂ embodiment in US-China trade," *Energy Policy*, vol. 34, issue 18, 2006, pp. 4063-4068.
- [7] A. Wyckoff and J. Roop, "The embodiment of carbon in imports of manufactured products: Implications for international agreements on greenhouse gas emissions," *Energy Policy*, vol. 22, issue 3, 1994, pp. 187-194.
- [8] G. P. Peters and E. G. Hertwich, "CO₂ embodied in international trade with implications for global climate policy," *Environmental science & technology*, vol. 42, issue 5, 2008, pp. 1401-1407.
- [9] I. Kolstad and A. Wiig, "What determines Chinese outward FDI?" *Journal of World Business*, vol. 47, issue 1, 2012, pp. 26-34.
- [10] Y. W. Cheung and X. W. Qian, "Empirics of China's outward direct investment," *Pacific Economic Review*, vol. 14, issue 3, 2009, pp. 312-341.
- [11] M. Habib and L. Zubawicki, "Corruption and foreign direct investment," *Journal of International Business Studies*, 33, 2002, pp. 291-307.
- [12] M. Aleksynska and O. Havrylchyk, "FDI from the south: The role of institutional distance and natural resources," *European Journal of Political Economy*, vol. 29, issue 284, 2013, pp. 38-53.
- [13] S. Lee and T. N. Sriram, "Sequential point estimation of parameters in a threshold AR(1) model," *Stochastic Processes and their Applications*, vol. 84, issue 2, 1999, pp. 343-355.
- [14] H. Tong, "Threshold models in time series analysis—30 years on," *Statistics and Its Interface*, vol. 4, issue 2, 2011, pp. 107-118.
- [15] M. Mohsen and B. S. Mohsen, "The threshold impact of fiscal and monetary policies on inflation: Threshold Model Approach," *Journal of Money and Economy*, vol. 10, issue 4, 2015.
- [16] C. L. Chen, "The threshold effects of RMB exchange rate fluctuations on imports and exports," *Journal of Financial Risk Management*, vol. 01, issue 02, 2012, pp. 15-20.
- [17] B. Hansen, "Threshold effects in Non-Dynamic panels: Estimation, testing, and inference," *Journal of Econometrics*, vol. 93, issue 2, 1999, pp. 345-368.
- [18] J. Johnston, "Econometric Methods," Second ed. New York: McGraw-Hill, 1972, pp. 267-291.